## MATS442 Stochastic simulation — Problems 3, 31.1.2020

The problems marked with **C** have a part which needs to be implemented in a computer programming language. Please return your implementation by email to  $\langle \text{santeri.j.karppinen}(\cdot at \cdot) jyu.fi \rangle$ .

- 1. Consider the following properties of min and max.
  - (a) Prove the following result: Let  $x, y \in A \subset \mathbb{R}$  and let  $f : A \to \mathbb{R}$  be non-decreasing. Then,  $f(\min\{x, y\}) = \min\{f(x), f(y)\}.$
  - (b) Suppose  $x, y \in A \subset \mathbb{R}$  and  $f : A \to \mathbb{R}$  is non-increasing. What is the expression for  $f(\min\{x, y\})$ ?
  - (c) What happens if you replace min with max in (a) and (b)?
- **2.C** (Construction of IS confidence intervals.)

Consider Example 4.16, with  $x_0 = 4$ , that is, consider the importance sampling estimator

$$I_{p,q}^{(n)}(\mathbf{1}(\cdot \ge 4)) = \frac{1}{n4\sqrt{2\pi}} \sum_{k=1}^{n} \exp\left(-\frac{Y_i^2}{2} + 4(Y_i - 4)\right),$$

where  $Y_i = \tilde{Y}_i + x_0$  where  $(\tilde{Y}_i) \stackrel{\text{i.i.d.}}{\sim} \text{Exp}(4)$  (cf. Example 2.2).

- (a) Implement a function which calculates the importance sampling estimator and the associated 95% confidence interval.
- (b) Use antithetic variable technique to reduce the variance of the estimator, and construct a confidence interval for the antithetic estimator.
- **3.C** Let  $m \ge 1$  and suppose  $p(1), \ldots, p(m) > 0$  with  $\sum_{i=1}^{m} p(i) = 1$ , that is,  $p(\cdot)$  defines a p.m.f. on  $\mathbb{X} = \{1, \ldots, m\}$ , and let  $f : \mathbb{X} \to \mathbb{R}$ .
  - (a) The given Julia function discrete\_from\_uniform implements the method of Theorem 2.3 (finds the K corresponding a U).
  - (b) Implement a stratified sampling approach for uniforms as in Example 5.10, and use the given Julia function to approximate

$$\mathbb{E}_p[f(X)] = \sum_{i=1}^m p(i)f(i).$$

- (c) Try your function with f(i) = i and  $p(i) \propto i$  for  $i \in \mathbb{X} = \{1, \dots, 10\}$ .
- (d) Try what happens with n = 55 (repeat the experiment a few times). Can you explain why?

(\* Optional extra): Think what is the *complexity* of your algorithm in (b). That is, what is the order of computer operations your algorithm needs to perform in terms of m and n? Try to design an algorithm which is O(m + n).

- 4.C Suppose that we are interested in estimating the mean and the second moment of the random variable Y, where  $Y \mid X \sim N(X, X^2)$  (that is, Y given X = x is a Gaussian with mean x and variance  $x^2$ ) and where  $X \sim \mathcal{U}(-1, 1)$ .
  - (a) Use Monte Carlo directly: Produce realisations of Y and compute the empirical mean and the second moment of Y.
  - (b) Use Rao-Blackwellisation and produce the same estimates of Y using only realisations  $(X_k) \stackrel{\text{i.i.d.}}{\sim} p$ .
    - (Hint: Compute first  $\mathbb{E}[Y \mid X = x]$  and  $\mathbb{E}[Y^2 \mid X = x]$ .)
  - (c) Compare the sample variance of the two estimators.

Complete the proof of Theorem 4.23 (ii): assuming  $q(x) = 0 \implies p(x) = 0$ ,  $\mathbb{E}_p[w(X)] < \infty$  and and  $\bar{\sigma}_{p,q}^2 := \mathbb{E}_p[w(X)\bar{f}^2(X)] < \infty$ , where  $\bar{f}(x) = f(x) - \mathbb{E}_p[f(X)]$ , show that 5.

$$v_{p,q}^{(n)} := \sum_{k=1}^{n} (W_k^{(n)})^2 \left[ f(Y_k) - \hat{I}_{p,q}^{(n)}(f) \right]^2 \qquad \text{satisfies} \qquad n v_{p,q}^{(n)} \xrightarrow{n \to \infty} \bar{\sigma}_{p,q}^2$$

Hints:

- (i) Recall that  $f(Y_k) \hat{I}_{p,q}^{(n)}(f) = \bar{f}(Y_k) \hat{I}_{p,q}^{(n)}(\bar{f})$ . (ii) Notice that (for *n* large enough such that  $\sum_{j=1}^n w_u(Y_j) > 0$ )

$$n(W_k^{(n)})^2 = \frac{\frac{1}{n}w_u^2(Y_k)}{\left(\frac{1}{n}\sum_{j=1}^n w_u(Y_j)\right)^2}.$$

(iii) Using these, show that

$$nv_{p,q}^{(n)} = \frac{\left(\frac{1}{n}\sum_{k=1}^{n}w_{u}^{2}(Y_{k})\bar{f}^{2}(Y_{k})\right) + R(n)}{\left(\frac{1}{n}\sum_{j=1}^{n}w_{u}(Y_{j})\right)^{2}},$$

and observe that the term  $R(n) \to 0$  as  $n \to \infty$ .

- (iv) Conclude your proof by showing that the remaining expression (with R(n) removed) converges to  $\bar{\sigma}_{p,q}^2$ .
- Consider the following Fisher-Yates shuffle algorithm: Define the vector a := (1, 2, ..., n). 6. Then, for  $k = n, k = n - 1, \ldots, k = 2$  repeat
  - (i) Pick an independent  $J_k \sim \mathcal{U}(\{1, \ldots, k\})$ .
  - (ii) Exchange the elements  $a_k \leftrightarrow a_{J_k}$ .

Report the final vector a as a random permutation of  $\{1, \ldots, n\}$ , that is, each permutation is equally likely.

- (a) How can you transform  $U \sim \mathcal{U}(0, 1)$  into  $j \sim \mathcal{U}(\{1, \dots, k\})$  efficiently?
- (b) Show that the vector a after applying the above algorithm is a random permutation of  $\{1, ..., n\}$ .

(Hint: Start by computing the distribution of  $a_n$ , then  $a_{n-1} \mid a_n, \ldots$  and finally  $a_1 \mid a_2, \ldots, a_n$ .)