

## MATS442 Stochastic simulation — Problems 1, 17.1.2020

The problems marked with **C** have a part which needs to be implemented in a computer programming language. Please return your implementation by email to `<santeri.j.karppinen(at)jyu.fi>`.

1. Consider the rain drops example: Assume  $(H_n)_{n \geq 1} \stackrel{\text{i.i.d.}}{\sim} \text{Bernoulli}(\pi/4)$ , and suppose that we estimate the value of  $\pi$  by computing Monte Carlo estimates  $I_p^{(n)}(f)$  with  $f(h) := 4h$ .
  - (a) Compute  $\text{Var}_p(f(X))$ . (Hint: Remember that  $\text{Var}(Y) = \mathbb{E}[Y^2] - (\mathbb{E}Y)^2$ .)
  - (b) Determine how big  $n$  must be chosen in order to attain an error less than 0.1 with probability at least 95%. Do this both with
    - i) The asymptotic error result.
    - ii) The bound from Chebychev's inequality.
  - (c) How does  $n$  change above if the error must be less than 0.01 (keeping the same probability  $\geq 95\%$ ).
  - (d) (\* optional) How does  $n$  change if the error must be less than 0.1 but the probability is  $\geq 99\%$ ?
  - (e) (\*) What about if you use the Hoeffding inequality?
- 2.C Consider the rain drops example as in Problem 1.

- (a) Write a function which simulates  $n$  independent  $\text{Bernoulli}(\pi/4)$  random variables and calculates an asymptotic confidence interval for the average. Your function should be defined as

```
using Distributions, Statistics
function est_pi(n, pr)
    # Simulate, calculate mean m and confidence interval width d
    (est=m, tol=d)
end
```

where  $n$  is the number of samples and  $pr$  is the asymptotic probability of the confidence interval (like 0.95 for a 95% CI). The output `est` should contain your estimator (mean) and `tol` the confidence interval width, so that the CI is `est±tol`. (Hint: You may use `rand(Bernoulli(p), n)` to simulate  $n$  samples from  $\text{Bernoulli}(p)$ , `mean` to calculate the average and `var` for the sample variance (or `std` for the standard deviation). The standard normal quantile function  $\Phi(t)$  is `quantile(Normal(0,1), t)`.)

- (b) Use your function to calculate the estimator and 95% confidence interval with 100 samples.
  - (c) Calculate 1000 estimators and related confidence intervals, and check how often the true value was within your confidence interval.
3. Show that if  $U_1, U_2 \stackrel{\text{i.i.d.}}{\sim} \mathcal{U}(0, 1)$ , then the random vector  $(R, T)$  has a density

$$p_{R,T}(r, t) = \begin{cases} \frac{1}{2\pi} r e^{-r^2/2}, & 0 < t < 2\pi, 0 < r < \infty, \\ 0, & \text{otherwise,} \end{cases}$$

where  $R = \sqrt{-2 \log U_1}$  and  $T = 2\pi U_2$ .

(Hint: Observe first that  $R$  and  $T$  are independent.)

- 4.C Let  $p$  be the Cauchy p.d.f., that is, for  $x \in \mathbb{R}$ ,

$$p(x) = \frac{1}{\pi(1+x^2)}.$$

- (a) Find an algorithm which generates  $X \sim p$  using  $U \sim \mathcal{U}(0, 1)$ .  
 (b) Implement your method and check that you got it right by inspecting the histogram of 10000 simulated values.

(Hint: If  $\mathbf{x}$  is a vector of simulated values, you may write

```
using Plots # Install by using Pkg; Pkg.add("Plots")
h = histogram(x, normalize=true, bins=LinRange(-4,4,20))
p(t) = 1/(pi*(1+t^2)) # same as function p(t) 1/(pi*(1+t^2)) end
t = LinRange(-4, 4, 500)
plot!(h, t, p.(t))
```

which shows the histogram within  $[-4, 4]$  normalised so that it matches the density.)

5. Suppose  $p : \mathbb{R} \rightarrow (0, \infty)$  is a continuous p.d.f., and let  $F(x) = \int_{-\infty}^x p(y)dy$  denote the corresponding c.d.f..

Let  $U \sim \mathcal{U}(0, 1)$  and  $-\infty < a < b < \infty$ . Find the distribution of

- (a)  $Y_b := F^{-1}(F(b)U)$ ,  
 (b)  $Y_{a,b} := F^{-1}(F(a)(1 - U) + F(b)U)$ , and  
 (c)  $Y_a := F^{-1}(F(a) + (1 - F(a))U)$ .

(Hint: Find first the distribution of  $Z_b := F(b)U$ ,  $Z_{a,b} := F(a)(1 - U) + F(b)U$  and  $Z_a := F(a) + (1 - F(a))U$ , and draw a picture.)

- 6.C The early PRNGs were mostly linear congruential generators, that is, based on a recursion of the form

$$Z_n = (aZ_{n-1} + c) \pmod{M},$$

with some parameters  $a, c, M \in \mathbb{N}$  with  $a, c < M$  and a seed  $Z_0 < M$ . Then,  $U_n := Z_n/M$  were “random uniform  $\mathcal{U}(0, 1)$ .”

Implement the above PRNG with  $M = 2^{31}$ ,  $a = 2^{16} + 3$  and  $c = 0$  and any  $Z_0 \in [0, M)$  you want.

- (a) Plot 100 samples produced by your generator (`plot(u)`), and inspect visually whether the produced samples seem i.i.d.  $\mathcal{U}(0, 1)$ .  
 (b) Produce 3000 uniform samples from a unit cube with your generator, that is,  $V_k = (U_{3k+1}, U_{3k+2}, U_{3k+3})$ , and plot them and inspect the randomness as above.  
 (Hint: Produce 9000 points from your generator to the vector  $\mathbf{u}$  and you can plot them by:

```
using Plots; plotly()
scatter(u[1:3:end], u[2:3:end], u[3:3:end], markersize=1)
```

You can rotate the figure with mouse.)